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Risk of Type 1 Diabetes in the Offspring Born Through Elective or Non-elective Caesarean Section in Comparison to Vaginal Delivery: A Meta-analysis of Observational Studies

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Abstract

Background: Caesarean section (CS) has been associated with an increased risk of Type 1 Diabetes (T1D). The lack of exposure to maternal vaginal and anal microbiome and bypassing the labor process often observed in elective CS may affect neonatal immune system development. This study aims to summarize the effects of elective and non-elective CS on T1D risk in the offspring.

Methods: A systematic literature search was conducted online for publications providing data on elective and non-elective CS with T1D diagnosis in children and young adults, followed by a

meta-analysis from selected studies. Newcastle-Ottawa Scale and GRADEpro tool were applied for quality analysis.

Results: Nine observational studies comprising over 5 million individuals fulfilled the inclusion criteria. Crude OR estimates showed a 12% increased T1D risk from elective CS compared to vaginal delivery with significant heterogeneity. Adjusted ORs from seven studies did not show T1D risk differences from either CS category, and heterogeneity was detected between studies. Separate analysis of cohort and case-control studies reduced the heterogeneity and revealed a slight increase in T1D risk associated with elective CS in cohort studies (adjusted OR = 1.12 (1.01 - 1.24)), and a higher increased risk associated with non-elective CS in case-control studies (adjusted OR = 1.19 (1.06 - 1.34)).

Conclusion: Summarized crude risk estimates showed a small increased T1D risk in children and young adults born through elective CS compared to vaginal delivery, but with significant heterogeneity. Adjusted risk estimates by study design indicated a slightly increased T1D risks associated with elective or non-elective CS.

Keywords: Autoimmune disease, Caesarean section, Elective Caesarean section, Meta-analysis, Type 1 Diabetes.

Introduction

Risks and benefits of Caesarean section (CS) have never been more relevant as worldwide CS rates are at their highest [1]. Between 1990 and 2014 the estimated global average rates from 121 countries have climbed from 6.7% to 19.1%, with Europe experiencing a 13.8% increase to 25% and Northern America a 10% increase to 32.3%. This trend is expected to continue, especially in middle and high-income countries [1].

CS is associated with short and long-term risks for the child [2]. The procedure has been associated with higher risks of diseases related to the offspring's immune system, namely asthma and Type 1 Diabetes (T1D) [3-5, 2, 6]. Type 1 Diabetes incidence in children under 15 years old are increasing in most parts of the world, and rates are particularly high in several European and North American countries, as well as Australia [7]. In Europe the incidence is estimated at 15,000 in 2005 and projected to increase more than 60% to over 24,000 cases by 2020 [8]. A recent study on incidence rates from 1989 to 2013 at multiple centres in 22 European countries showed no clear indication of the rates slowing down in the majority of these countries [9]. The incidence rate in young adults and adults are lower but may be rising as well [10]. The disease results from autoimmune-destruction of pancreatic beta cells leading to inadequate insulin production. The disease can manifest at any age, but new diagnoses seem to peak in early adolescence [7]. A meta-analysis published in 2008 of more than 2 million CS-born children linked CS to an approximately 20% higher risk of developing T1D, after adjustment of maternal diabetes history and other possible confounders [4].

Bypassing the birth canal in CS is thought to reduce exposure to the mother's vaginal and anal microbiome, which may result in subsequent suboptimal development of the child's immune system. The so-called "hygiene hypothesis" proposes that certain microbiome colonization is necessary for healthy immune system maturation, and the lack of its diversity and numbers have been associated with the development of immune disorders including T1D [11-15]. Studies have described the different microbiome colonization apparent in children born through CS compared to vaginal delivery [16, 11] and children with or without T1D [17]. Observed rising autoimmune disease frequencies, including T1D, and lower infection rates due to more hygienic living environment support the hypothesis [18].

Another difference between vaginal delivery (VD) and CS that may also contribute to developing T1D is the possibility to bypass "labor". This process is responsible for the HPA (Hypothalamic Pituitary Axis) activation, which initiates a cascade of reactions resulting from "labor stress" that prepares the neonate to adapt to its new environment outside the mother [19]. Omission of this process could be followed by delayed risks of immune development impairments [14, 20]. As more women deliver by CS even without medical indication, as reported in a U.S. study [21], the potential disadvantage of pre-labor birth becomes

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increasingly pertinent. Moreover, it is possible that both theories describe mechanisms through which elective CS, in particular, may contribute to the development of T1D [14, 22]. Extensive studies relating CS to T1D have been published, but in contrast to the previous metaanalysis [4], we focus particularly on studies that distinguish between the different types of CS. This paper aims to systematically review studies that highlight elective (planned) and nonelective CS associated with T1D development in children and young adults and produce a risk estimate from each procedure compared to vaginal delivery. It also provides an analysis of the association between CS, in general, as compared with VD and T1D risk from the selected studies.

Methods

Search strategy and study selection

A systematic search strategy was developed following the PICO (Patient, Intervention, Comparison, Outcome) framework. Applied search terms were *(((birth OR delivery OR caesarean OR cesarean OR labor)) AND (child OR offspring OR adolescent OR young adult)) AND (Type 1 Diabetes OR IDDM OR T1D)* with limits to humans. The searches were conducted on MEDLINE, Web of Science and CINAHL (The Cumulative Index to Nursing and Allied Health Literature) databases for articles published before 23rd April 2018. Duplicates were removed prior to screening articles by their title or abstract. These articles were then subject to full-text screening for eligibility. This search and study selection process were performed by two independent reviewers (JT and AG) and any discordance resolved by discussion with a third reviewer (HB).

Eligibility and exclusion criteria

Original studies that reported effect sizes or adequate data for calculation of T1D risk in children or young adults associated with elective CS and non-elective CS on were included. Diagnosis was restricted up to the age of 24 years, applying the "young adult" definition by the World Health Organization [23]. When studies indicated availability of data, authors were contacted to provide additional effect size estimates. Studies which only reported CS without specifying whether elective or non-elective (emergency) were excluded.

Data extraction and statistical analysis

Data extraction included the number of children born by each mode of delivery, the number of cases and controls from case-control studies or the number of children who were and were not diagnosed with T1D by the end of the study period from cohort studies, and confounding variables. These study characteristics were then presented in a table. Available adjusted hazard ratios (HR), risk ratios (RR) and odds ratios (OR), as well as, the adjusted confounders were either extracted from the papers, or calculated when the data were available (using the Mantel-Haenszel method), or gathered from contacted authors. As T1D is considered a rare disease, proportional HR, RR and OR are treated similarly [24] and presented as OR.

Summary crude OR for risk of CS compared to VD on T1D were estimated by combining data from the previous meta-analysis [4] and more studies identified in this review. Similar studies were only included once for the overall effect estimate. Risk of T1D estimates associated with elective CS compared to VD and emergency CS compared to VD were calculated only from studies eligible for this review. In addition, elective CS with non-elective CS were also briefly compared.

Number of cases and non-cases in each mode of delivery were applied in the pooled crude OR estimates. Adjusted ORs and corresponding standard errors (SE) calculated from the 95% confidence interval (CI) were applied in the pooled adjusted effect size estimates (aOR).

Inverse weighted meta-analysis using a random effects model for binary data was performed to estimate pooled effect sizes. Heterogeneity estimates, I^2 , indicate how much variability in the estimates is due to heterogeneity, and τ^2 (investigated with DerSimonian and Laird method) indicates the total amount of heterogeneity [25, 26]. Funnel plots were generated to check for publication bias. Subgroup analyses divided the studies into cohort and case-control studies were performed to identify whether the difference in study design is caused by the different study designs. All statistical analyses were performed with R software (version 3.3.2). Estimates are considered significant when p value is <0.05.

Study and evidence quality assessment

Individual study quality was determined using the Newcastle-Ottawa Scale (NOS) for observational studies [27]. The overall quality of evidence was assessed using the GRADE guidelines [28-30].

Results

Systematic literature search

Literature database search yielded 4,376 publications, including review articles and meeting proceedings. After deduplication, we screened the titles and abstracts and removed 3,481 articles. The full texts of the 104 remaining articles were appraised following the inclusion criteria. Eighty-nine articles did not report data separating elective and non-elective CS, two reported part of or the same studies, and one reported a population within a geographical scope and time period of a larger study. One study was excluded as it investigated only children with high-risk HLA genotype, leaving eleven studies to be potentially included in our meta-analysis. Four of the eleven studies only presented crude effect sizes and one reported an adjusted effect estimate for elective CS and VD but not for non-elective CS and VD. Adjusted effect sizes for both types of CS and VD were reported in three studies. In one study sufficient information about maternal diabetes enabled a calculation of effect size adjusted for that variable using the Mantel-Haenszel method. Authors of six articles were contacted for unreported adjusted effect sizes, another one for data on elective and non-elective CS. Three responded with the necessary information, while the latter was unable to provide the data. One author did not respond to a request to clarify their definition of primary CS regarding electiveness, and the article was excluded. Finally, nine studies were included in the quality and meta-analysis. The selection process is illustrated in ESM_Figure 1_Study selection.

Study characteristics

There were nine observational studies, five cohort and four case-control studies, included in this analysis. Eight were conducted in Europe and one in Australia. Most studies ascertained cases from regional or national patient registries, and one used hospitalization records. Two studies from Scotland have a short overlapping time period in one region (Patterson, et al [31] and Robertson, et al [32]), and two studies from Sweden also utilised data from an overlapping time period but in different study designs (Khashan, et al [33] and Samuelsson, et al [34]). The age of diagnosis differs between studies, with one cohort study following children from birth until the age of 5- 6 years, while others defined their cases until age 15, 18 and the oldest at 27 years old.

Having extracted their data from official registries, most cohort studies are able to adjust for some confounding variables deemed relevant to T1D risk. Maternal diabetes is a common confounding variable, as well as gestational age and birth weight. One study also considered paternal age and history of diabetes. Another study presented an analysis which includes CS and T1D data in siblings to adjust effect estimates. Relevant study characteristics are presented in Table 1.

Meta-analysis

Before distinguishing between elective and non-elective CS we present an overall CS result. The basis for this analysis is studies from the previous meta-analysis by Cardwell et al (20 studies) [4], in which one of the studies distinguished between elective and emergency CS [35], while another mentioned finding a significant difference for elective CS as an additional analysis [31]. The seven later studies that made this distinction were then added. Table 2 shows a summary of crude and adjusted risk estimates between all CS and VD, elective CS and VD, non-elective CS and VD and between elective and non-elective CS from all studies. Four studies differentiated spontaneous VD with assisted or instrumental VD. For the purpose of this analysis, the records from both procedures are merged as VD in crude risk summaries, and reported adjusted risk estimates from only spontaneous VD is applied to adjusted risk summary.

A summarized crude OR for CS compared to VD on the T1D risk 27 studies showed that there was a significant increased risk (OR 1.12 (1.05 - 1.20)). However, significant heterogeneity was detected (Table 2). Analysis focused on the nine currently included studies showed an insignificant overall crude OR of 1.06 (0.98 - 1.15) for CS on T1D risk with significant heterogeneity across studies. Separate analysis from cohort studies did not show any significant risk, but there was significant summary OR of 1.25 (1.03 - 1.52) from case-control studies with insignificant heterogeneity.

Figure 1 shows the risk estimates for T1D in elective CS compared to VD from nine studies identified in the systematic review. A significant higher risk (OR 1.12 (1.00 - 1.24) was

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estimated, with significant heterogeneity. Analysis on separate study designs did not indicate significant risk estimates from both groups, and significant heterogeneity estimates remained. The risk of non-elective CS on T1D compared to VD (Figure 1) according to the summary crude OR was 1.05 (0.89 - 1.24). Subgroup results from cohort and case-control studies showed insignificant heterogeneity within each group, but the crude ORs of 0.95 (0.88 - 1.02) and 1.33 (1.07 - 1.65), respectively, indicated a higher T1D risk posed by non-elective CS in case-control studies.

Five cohort and two case-control studies enabled estimations from adjusted ORs (aOR) of elective CS compared to VD and non-elective CS compared to VD. Adjustment variables vary between studies as presented in Table 1.

The forest plots in Figure 2 (upper plot) shows an overall aOR for elective CS on T1D risk from these seven studies was 1.09 (0.97 - 1.22). Focusing on only cohort studies showed an aOR of 1.12 (1.01 - 1.24) and no significant heterogeneity was found. The aOR of 1.14 (0.71 - 1.85) from the case-control studies was similar in magnitude but not significant, and showed considerable heterogeneity although based on only two studies.

As for non-elective CS, the overall aOR was insignificant (Figure 2, lower plot). The cohort study summarized aOR of 1.00 (0.92, 1.09) was not significant but the case-control study aOR of 1.19 (1.06 - 1.34) was significant with little heterogeneity observed.

Publication bias analysis

Funnel plots of studies on the risk of elective CS or non-elective CS compared to VD showed slight asymmetries, particularly for the elective CS analysis (ESM_Figure 2_Funnel plots). However, most of the studies investigated CS as one factor among others suspected to be

associated with T1D risk, reducing the likelihood of publication bias due to positive findings. Moreover, as there are only a small number of studies, the plots are difficult to interpret and any evidence for publication bias is weak.

Qualitative analysis

Newcastle-Ottawa scales (NOS) for cohort and case-control studies were applied to determine the study quality. The cohort studies scored 7-9 out of 9, and the case-control studies scored 8-9 out of 9 indicating high quality observational studies. Detailed assessment is available as supplementary materials (ESM_Table 1). This result is summarized and applied into the GRADE risk of bias assessment as not serious.

Following the GRADE study quality assessment criteria, overall assessment of risk of bias and indirectness were not considered serious, although there were concerns about inconsistency, imprecision and publication bias. Taken together, the assessment concluded that there was very low certainty of the evidence, mainly due to high heterogeneity between the small number of studies and that the summarized ORs did not show a significant difference between delivery methods (summary of findings presented in Supplementary materials ESM Table 2).

Discussion

This systematic review found nine studies that highlighted the association between elective and non-elective CS procedures in comparison to VD and T1D risk in children and young adults. This meta-analysis includes four case-control studies with 10,925 cases of T1D and 39,543 controls in total and five cohort studies with 16,868 cases of T1D in a total cohort size of 5,261,891 individuals. To our best knowledge, this is the first meta-analysis that assesses the risks of T1D in elective and non-elective CS compared to VD. Our meta-analysis of crude risk estimates found a significant increased T1D risk associated with elective CS, but there is significant heterogeneity between studies. Adjusted risk assessment of all studies did not show any significant differences in risk estimates from elective or non-elective CS compared to VD, and heterogeneity was detected. Separate analyses between study-types revealed that results from the cohort studies showed an aggregated 12% higher risk of T1D in the offspring born through elective CS compared to VD, while case-control studies showed a 19% increase in the risk in the offspring born through nonelective CS. Heterogeneity within these subgroups were insignificant.

The findings in this study adds to the knowledge gained in a previous meta-analysis where a 20% increase of T1D risk due to CS was estimated [4]. The excess risk estimate from nine studies (unadjusted) and from five cohort studies (adjusted) in this meta-analysis are in line with a hypothesis first advanced in the study by Patterson, et al, which was also included in the previous meta-analysis but was not further investigated [31]. A higher risk from elective CS compared to VD on T1D risk in the offspring supports both the hygiene and labor-stress theories, although the exact mechanisms of the causal pathways are still unclear [20]. Moreover, elective CS is often performed in the early term period and this may expose the offspring to premature birth. One study found reduced T1D risk with each additional week after 39 weeks of gestation [36]. Another study found that early-term deliveries (gestational week 37 to 38+6 days) increases the risk of T1D diagnosed at age 5-18 years by approximately 50%. The study stipulated that the "lost" days put the newborns at a similar risk as the preterm births [37]. Moreover, aberrant immune cell maturity have been observed in premature children born through pre-labor CS [38].

All the publications in this meta-analysis are regarded as high-quality observational studies (scoring 7-9 out of 9), as determined by our application of NOS. However, the scales have been

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criticised due to varied inter-rater reliability [39]. In our case, we found a relatively homogeneous quality of all studies included in our meta-analysis, so the judgement did not affect the final results. We believe the scale has some benefit despite its limitations; therefore, we decided to use it. In addition, the GRADE tool provided an additional assessment to evidence quality from the collected studies. The assessment resulted in "low quality of evidence" grade since only observational studies are included. Significant heterogeneity was also observed but this was reduced after subgroup analyses. Therefore, the discordance in which category of CS is associated with increased risk of T1D (elective CS in the cohort studies and non-elective CS in the case-control studies) is not likely explained by the study design. All adjusted risk estimates controlled for maternal diabetes, which is a known major risk factor in T1D. However, there are other potentially important covariates not available in the self-calculated adjusted risk estimate from the Samuelsson, et al [34] case-control study, such as maternal age, birth weight and gestational age.

Without having similarly-adjusted estimates across all studies, the possibility that adjustments might affect the slight increased risks from both types of CS seen in this meta-analysis cannot be ruled out. Paternal diabetes is a known covariate for T1D in the offspring [40, 41]. One of the studies in this analysis also found higher T1D risk associated with paternal history of the disease in their multivariate model analysis [42]. Another potentially important covariate that was not available in most studies here is paternal age. A large study in Northern Ireland found an approximately 50% increase of T1D can be contributed to fathers aging >35 years [36]. Maternal BMI was also not commonly adjusted for in these studies although there is evidence that high maternal BMI in the first trimester increased T1D risk by 20-48% independent of parental diabetes [43].

Birth weight for gestational age, instead of birth weight and gestational age separately, may be a better adjustment variable as it considers maturity rather than growth alone, as small-forgestational-age is found to be a protective factor for T1D [44]. Admittedly, adjustment of both birthweight and gestational age may act as sufficient substitutes. Higher birth order (second or later) seem to have a protective effect, particularly in early onset T1D (<5 years old) [45]. Among the studies in this systematic review, two included being first born or later in their analysis [42, 33]. Adjusting analysis with data on siblings born by a different delivery mode and did not develop T1D may also be a useful approach to control for unmeasured familial traits such as shared genes and lifestyle. Khashan, et al found that elective CS slightly increases T1D after adjusting for confounders such as maternal diabetes, maternal age, and gestational age, but controlling for siblings eliminated the significance of the previous finding, suggesting that unspecified familial traits accounted for the association [33].

Limitations

Studies included in this analysis are from developed countries in Europe and Australia. All efforts have been made to identify and include all relevant studies, but relatively few studies met the selection criteria. Data unavailability on specific types of delivery mode may be a reason for lack of reports from other countries.

Four of the studies in two countries utilized data from registers that might overlap with each other in certain time periods. However, each pair was either a cohort or case-control study, and separating the study types reduced the possibility of an exaggerated association. Moreover, the case-control study of the pair did not provide any adjusted effect estimate, and is excluded from the adjusted risk calculation.

An attempt to isolate the presence of labor by way of comparing CS with or without labor was inconclusive due to insufficient information. Only two studies reported crude risk estimates from in-labor CS [44, 42]. Moreover, the non-elective CS data presented in the studies were not always categorized as in-labor CS, and the traditional classification of all unscheduled CS as emergency CS does not necessarily indicate failed labor [46, 47].

It should be emphasized that any role of CS in the development of T1D, is likely the result of a combination of genetic, prenatal, perinatal, environmental, and lifestyle factors [48]. Immune system development begins early in fetal life and may show aberrations (i.e. due to HLA class II or PTPN22 gene polymorphisms known to confer susceptibility to T1D) seen in T1D preceding any type of delivery [49]. However, interaction between CS and immune response genes IFIH1 and CD25 has also been associated with the manifestation of T1D [50]. Breastfeeding, time of introduction to cow's milk or gluten-containing-foods, and childhood infections have all been observed to influence the progression from islet antibodies to T1D [51-55, 41]. Although study results on many non-hereditary T1D risk factors remain inconsistent, further research into intrauterine and postnatal immune development is essential to confirm or refute these observations, especially regarding early onset T1D cases where it is more plausible that these factors play a role [41].

Conclusion

Despite existing theories on the roles of hygiene and labor-stress in the development of autoimmunity in T1D, few studies report elective and non-elective CS as separate risk factors. A systematic literature search identified nine studies that investigated the association of these delivery modes to T1D risk in the offspring. All studies were observational, with five cohorts having follow-up periods ranging from birth until 6 to 24 years of age, and four case-control studies that include T1D cases until 15 years of age. Summarized crude and adjusted risk estimates from all studies found an increased T1D risk associated with elective CS, but there was significant heterogeneity among studies. This was reduced after the studies were grouped into cohort and case-control studies, and the adjusted risk estimates indicated a slightly increased T1D risk related to elective or non-elective CS depending on the study design. Discrepancies in results from subgroup analyses may be due to the inconsistent adjustment of covariate variables between studies, except for maternal diabetes. A pooled study analysis with more studies and similarly adjusted variables may shed better light on this possibility. Until then, this meta-analysis indicated that T1D risk might be related to modes of delivery.

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Compliance with Ethical Standards

Conflict of Interest

Justine Tanoey, Amit Gulati, Chris Patterson, and Heiko Becher declare that they have no conflict of interest.

Human and Animal Rights and Informed Consent

This article does not contain any studies with human or animal subjects performed by any of the authors.

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| Author, country & | Study | Casa diagnosis | Age at diagnosis; | Vaginal Delivery | | Elective C-Section | | Non-Elective C-Section | | Adjustment variables | |
|--|---|--|---|------------------|------------------------|-----------------------|------------------------|---------------------------|------------------------|--|--|
| publication year | publication design year Scotland hosp discharge database ar | Case diagnosis | length of follow-up | Cases | Controls/ Non-cases | Cases | Controls/ Non-cases | Cases | Controls/ Non-cases | Aujustment variables | |
| Patterson, et al, UK (Scotland), 1994 [31] | C-C | Scotland hospital discharge database and Regional Diabetes registry (Scottish Study Group for the Care of Diabetes in the Young) | Onset < 15 yo; 1975 - 1976 | 236 | 1,241 | 24 | 60 | 10 | 52 | Maternal diabetes Maternal age Parity Gestational age Birthweight Social class | |
| McKinney, et al, UK (Yorkshire), 1997 [35] | C-C | Regional Diabetes registry | Onset < 16 yo; 1993 - 1994 | 162 | 290 | 16 | 15 | 18 | 20 | - | |
| Algert, et al, Australia, 2009 [44] | Cohort | Hospitalization | Onset < 6 yo; 2000 – 2005/ FUª: up to 6 years | 190 | 377,078 | 46 | 69,485 | 36 | 55,205 | Maternal diabetes Preeclampsia SGA (<10th centile) LGA (>90th centile) Preterm birth | |
| Robertson and Harrild, UK | C-C | Regional Diabetes registry | Onset < 15 yo; 1984 - 2002 | 308 | 940 | 24 | 53 | 29 | 90 | - | |

Table 1. Study characteristics (continued on the next 3 pages).

| (Grampian, Scotland), 2010 [32] | | (Scottish Study Group for the Care of Diabetes in the Young) | | | | | | | | |
|--|--------|--|---|-------|-----------|-----|---------|-----|---------|--|
| Khashan, et al, Sweden, 2014 [33] | Cohort | National registry (Swedish National Patient Register) | Onset < 15 yo; 1982 – 2009/ FUª: up to 27 years | 9 025 | 2,277,914 | 678 | 158,820 | 725 | 190,921 | Offspring age as a time dependent variable Year of birth Gestational age Maternal pre- pregnancy diabetes^b Birth order^b Maternal age^b Body Mass Index^b Country of birth^b Education^b Gestational diabetes^b Small/Large for Gestation^b Preeclampsia^b |
| Black, et al (1), UK, (Scotland) 2015 [56] | Cohort | Regional Diabetes registry (Scottish Care Information Diabetes Collaboration – SCI DC) | Not specified; 1993 – 2007/ FUª: up to 21 years | 1,260 | 251,657 | 82 | 12,273 | 250 | 55,765 | Maternal age Gestation at birth Carstairs deprivation score Maternal smoking Birthweight Year of delivery Infant gender Breastfeeding at 6 weeks |

| | | | | | | | | | | • Maternal Type 1 Diabetes (T1D) |
|--|--------|--|---|-------|-----------|-----|---------|-----|---------|---|
| Clausen, et al, Denmark, 2015 [57] | Cohort | National registry | Onset < 15 yo; 1982 – 2010/ FUª: up to 15 years | 3,762 | 1,493,850 | 302 | 122,487 | 336 | 139,599 | Year of birth Parental age at childbirth Parental education level Parental T1D First or higher birth order Gestational age^c Birth weight^c |
| Samuelsson, et al, Sweden, 2015 [34] | C-C | National Diabetes study registry (Swedish paediatric diabetes quality register - SWEDIABKIDS) | Onset < 19 yo; 1984 – 2012 | 7,999 | 32,530 | 924 | 3,720 | 453 | 1,254 | Maternal diabetes |
| Black, et al (2), UK, (Scotland) 2016 [58] | Cohort | Regional Diabetes registry (Scottish Care Information Diabetes Collaboration – SCI DC) | Onset < 21 yo; 1993 – 2007/ FUª: up to 21 years | 68 | 13,311 | 75 | 17,844 | 33 | 8,814 | Maternal age Gestation at birth Carstairs deprivation score Maternal smoking Birthweight Year of delivery Infant gender Breastfeeding at 6 weeks |

• Maternal T1D

^a Follow-up.

^b Adjustment for these additional variables to the reported risk estimates did not change results.

^c Restricting analyses to term births (>37 weeks gestation) or birth weight >2500g did not change results.

| | | All CSª com OR (| pared to VD ^a 95%CI) | Elective CSª co OR (| ompared to VD ^a 95%Cl) | Non-El compa OR (| ective CS ^a red to VD ^a 95%Cl) |
|--|-----------------------|------------------------------------|-------------------------------------|-------------------------|--------------------------------------|-------------------------|--|
| Included studies | | Random | Heterogeneity | Random | Heterogeneity | Random | Heterogeneity |
| Cardwell, et al [4] | Crude | 1.23 (1.15 – 1.32) | $I^2 = 0\%$ p = 0.54 | NA | NA | NA | NA |
| | Adjusted | 1.19 (1.04 – 1.36) ^b | l ² = 0% p = 0.69 | NA | NA | NA | NA |
| Cardwell, et al [4] + 7 studies (N = 27) | Crude ^c | 1.12 (1.05 – 1.20) | l ² = 49.8% p = 0.002 | NA | NA | NA | NA |
| 9 studies | Crude | 1.06 (0.98 – 1.15) | l ² = 68.1% p = 0.002 | 1.12 (1.00 – 1.24) | $l^2 = 64.4\%$ p = 0.004 | 1.05 (0.89 – 1.24) | l ² = 85.4% p < 0.01 |
| 7 studies ^d | Adjusted ^e | NA | | 1.09 (0.97 – 1.22) | l ² = 72.7% p < 0.01 | 1.04 (0.94 – 1.14) | l ² = 51.9% p = 0.05 |

Table 2. Crude OR and heterogeneity of all CS^a compared to VD^a, elective CS^a compared to VD^a, and non-elective CS^a compared to VD^a.

^a CS = Caesarean section, VD = Vaginal delivery,

^b Adjusted for birthweight, gestational age, maternal age, birth order, breastfeeding, and maternal diabetes.

^c Adjusted estimates were not available.

^d Adjusted estimates from McKinney, et al and Robertson and Harrild, et al studies were not available.

^e Adjustments in each study according to the variables listed in the study characteristics (Table 1).

Figure 1 Crude OR estimates of elective (upper) and non-elective (lower) Caesarean Section (CS; N – Total) risk on offspring Type 1 Diabetes (T1D) compared to vaginal delivery (VD; N – Total).

| | Crude (| OR estim | ates c | omparing E | lective CS and VD | | | |
|--|--|---|---|--|------------------------------------|--|---|--|
| | Elect | ive CS | 1 | /D | | | | |
| Author & year of study / Study design | T1D | N | T1D | N | Odds Ratio | OR | 95%-CI | Weight |
| Study design = case-control | | | | | | | | |
| Patterson, et al., 1994 | 24 | 84 | 236 | 1477 | | 2.10 | [1.28; 3.45] | 3.9% |
| McKinney, et al., 1997 | 16 | 31 | 162 | 452 | | 1.91 | [0.92; 3.96] | 2.0% |
| Robertson & Harrild, 2010 | 24 | 77 | 308 | 1248 | | 1.38 | [0.84; 2.28] | 3.8% |
| Samuelsson, et al., 2015 | 924 | 4644 | 7999 | 40529 | | 1.01 | [0.94; 1.09] | 22.2% |
| Random effects model | | 4836 | | 43706 | + | 1.44 | [0.95; 2.19] | 31.9% |
| Heterogeneity: $I^2 = 76\%$, $\tau^2 = 0.1257$, $p < 0.0$ |)1 | | | | | | | |
| Study design = cohort | | | | | | | | |
| Algert, et al., 2009 | 46 | 69531 | 190 | 377268 | | 1.31 | [0.95; 1.81] | 7.6% |
| Khasan, et al., 2014 | 678 | 159498 | 9025 | 2286939 | | 1.08 | [1.00; 1.17] | 22.1% |
| Black, et al., 2015 | 82 | 12355 | 1260 | 252917 | | 1.33 | [1.07; 1.67] | 11.9% |
| Black, et al., 2016 | 75 | 17919 | 68 | 13379 | | 0.82 | [0.59; 1.14] | 7.4% |
| Clausen, et al, 2016 | 302 | 122789 | 3762 | 1497612 | - | 0.98 | [0.87; 1.10] | 19.2% |
| Random effects model | | 382092 | | 4428115 | \$ | 1.08 | [0.96; 1.22] | 68.1% |
| Heterogeneity: $I^2 = 60\%$, $\tau^2 = 0.0099$, $p = 0.0099$ |)4 | | | | | | | |
| Random effects model | | 386928 | | 4471821 | - | 1.12 | [1.00; 1.24] | 100.0% |
| Heterogeneity: $I^2 = 64\%$, $\tau^2 = 0.0117$, $p < 0.020$ | 11 | | | | | | | |
| Residual heterogeneity: $I^2 = 69\%$, $p < 0.01$ | | | | | 0.5 1 2 | | | |
| | | | | | | | | |
| | | | | CONTRACTORS IN CONTRACTORS | | | | |
| <u>c</u> | rude OF | estimate | es con | nparing Nor | -Elective CS and VD | | | |
| <u>c</u> | Non-Ele | estimate | es com | nparing Nor | n-Elective CS and VD | 0.5 | 05% 01 | Malak |
| <u>C</u> Author & year <mark>of</mark> study / Study design | Non-Ele T1D | estimate ective CS N | es com T1D | nparing Nor VD N | Odds Ratio | OR | 95%-CI | Weight |
| <u>C</u> Author & year of study / Study design Study design = case-control | Non-Ele T1D | R estimate ective CS N | es com T1D | nparing Nor VD N | Odds Ratio | OR | 95%-CI | Weight |
| C Author & year of study / Study design Study design = case-control Patterson, et al., 1994 | rude OF Non-Ele T1D | R estimate ective CS N 62 | es com T1D 236 | nparing Nor VD N 1477 | Odds Ratio | OR 1.01 | 95%-Cl [0.51; 2.02] | Weight |
| C Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 | Non-Ele T1D | R estimate ective CS N 62 38 | es com T1D 236 162 | nparing Nor VD N 1477 452 | Odds Ratio | OR 1.01 1.61 | 95%-Cl [0.51; 2.02] [0.83; 3.13] | Weight 4.3% 4.6% |
| C Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 | rude OF Non-Ele T1D | ective CS N 62 38 119 | es com T1D 236 162 308 | nparing Nor VD N 1477 452 1248 | Delective CS and VD Odds Ratio | OR 1.01 1.61 0.98 | 95%-Cl [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] | Weight 4.3% 4.6% 7.9% |
| C Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 | 2rude OF Non-Ele T1D 10 18 29 453 | R estimate sective CS N 62 38 119 1707 | es com T1D 236 162 308 7999 | nparing Nor VD N 1477 452 1248 40529 | Delective CS and VD | OR 1.01 1.61 0.98 1.47 | 95%-Cl [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] | Weight 4.3% 4.6% 7.9% 16.3% |
| Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model | 2rude OF Non-Ele T1D 10 18 29 453 | R estimate ective CS N 62 38 119 1707 1926 | es com T1D 236 162 308 7999 | nparing Nor VD N 1477 452 1248 40529 43706 | Delective CS and VD Odds Ratio | OR 1.01 1.61 0.98 1.47 1.33 | 95%-CI [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] [1.07; 1.65] | 4.3% 4.6% 7.9% 16.3% 33.1% |
| Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity. $J^2 = 27\%$, $\tau^2 = 0.0165$, $p = 0.2$ | 10 10 18 29 453 | estimate ective CS N 62 38 119 1707 1926 | 236 162 308 7999 | nparing Nor VD N 1477 452 1248 40529 43706 | Delective CS and VD Odds Ratio | OR 1.01 1.61 0.98 1.47 1.33 | 95%-Cl [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] [1.07; 1.65] | 4.3% 4.6% 7.9% 16.3% 33.1% |
| Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $J^2 = 27\%$, $\tau^2 = 0.0165$, $p = 0.2$ Study design = cohort | 10 10 18 29 453 | ective CS N 62 38 119 1707 1926 | 236 162 308 7999 | nparing Nor VD N 1477 452 1248 40529 43706 | Delective CS and VD Odds Ratio | OR 1.01 1.61 0.98 1.47 1.33 | 95%-Cl [0.51; 2.02] [0.63; 3.13] [0.63; 1.52] [1.32 ; 1.64] [1.07; 1.65] | 4.3% 4.6% 7.9% 16.3% 33.1% |
| Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity. $J^2 = 27\%$, $\tau^2 = 0.0165$, $p = 0.2$ Study design = cohort Algert, et al., 2009 | rude OF Non-Ele T1D 10 18 29 453 5 36 | ective CS N 62 38 119 1707 1926 55241 | es com T1D 236 162 308 7999 190 | nparing Nor VD N 1477 452 1248 40529 43706 377268 | Deflective CS and VD Odds Ratio | OR 1.01 1.61 0.98 1.47 1.33 1.29 | 95%-Cl [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] [1.07; 1.65] [0.91; 1.85] | Weight 4.3% 4.6% 7.9% 16.3% 33.1% 9.7% |
| C Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harriid, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity. $J^2 = 27\%$, $\tau^2 = 0.0165$, $p = 0.2$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 | rude OF Non-Ele T1D 10 18 29 453 5 36 725 | € estimate ective CS N 62 38 119 1707 1926 55241 191646 | es com T1D 236 162 308 7999 190 9025 | N N 1477 452 1248 40529 43706 377268 2286939 | Delective CS and VD Odds Ratio | OR 1.01 1.61 0.98 1.47 1.33 1.29 0.96 | 95%-Cl [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] [1.07; 1.65] [0.91; 1.85] [0.89; 1.03] | Weight 4.3% 4.6% 7.9% 16.3% 33.1% 9.7% 16.9% |
| Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity. ² = 27%, τ^2 = 0.0165, p = 0.2 Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 Black, et al., 2015 | rude OF Non-Ele T1D 10 18 29 453 5 5 36 725 250 | R estimate sctive CS N 62 38 119 1707 1926 55241 191646 56015 | es com T1D 236 162 308 7999 190 9025 1260 | N N 1477 452 1248 40529 43706 377268 2286939 252917 | Delective CS and VD | OR 1.01 1.61 0.98 1.47 1.33 1.29 0.96 0.90 | 95%-Cl [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] [1.07; 1.65] [0.91; 1.85] [0.89; 1.03] [0.78; 1.03] | Weight 4.3% 4.6% 7.9% 16.3% 33.1% 9.7% 16.9% 15.7% |
| Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $J^2 = 27\%$, $\tau^2 = 0.0165$, $p = 0.2$ Study design = cohort Algert, et al., 2019 Khasan, et al., 2014 Black, et al., 2015 Black, et al., 2016 | rude OF Non-Ele T1D 10 18 29 453 5 5 36 725 250 33 | Cestimate ective CS N 62 38 119 1707 1926 55241 19164 56015 8847 | 236 162 308 7999 190 9025 1260 68 | N N 1477 452 1248 40529 43706 377268 2286939 252917 13379 | Delective CS and VD Odds Ratio | OR 1.01 1.61 0.98 1.47 1.33 1.29 0.96 0.90 0.73 | 95%-Cl [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] [1.07; 1.65] [0.91; 1.85] [0.89; 1.03] [0.78; 1.03] [0.78; 1.03] [0.48; 1.11] | Weight 4.3% 4.6% 7.9% 16.3% 33.1% 9.7% 16.9% 15.7% 8.3% |
| C Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $J^2 = 27\%$, $\tau^2 = 0.0165$, $p = 0.2$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 Black, et al., 2016 Clausen, et al., 2016 | rude OF Non-Ele T1D 10 18 29 453 5 5 36 725 250 33 333 | Cestimate ective CS N 62 38 119 1707 1926 55241 191646 56015 8847 139935 | 236 162 308 7999 190 9025 1260 68 3762 | N N 1477 452 1248 40529 43706 377268 2286939 252917 13379 1497612 | Delective CS and VD Odds Ratio | OR 1.01 1.61 0.98 1.47 1.33 1.29 0.96 0.90 0.73 0.96 | 95%-Cl [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] [1.07; 1.65] [0.91; 1.85] [0.89; 1.03] [0.78; 1.03] [0.48; 1.11] [0.85; 1.07] | Weight 4.3% 4.6% 7.9% 16.3% 33.1% 9.7% 16.9% 15.7% 8.3% 16.3% |
| C Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity. $J^2 = 27\%$, $\tau^2 = 0.0165$, $p = 0.2$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 Black, et al., 2016 Clausen, et al., 2016 Random effects model | rude OF Non-Ele T1D 10 18 29 453 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 | Cestimate ective CS N 62 38 119 1707 1926 55241 191646 56015 8847 139935 451684 | 236 162 308 7999 190 9025 1260 68 3762 | N N 1477 452 1248 40529 43706 377268 2286939 252917 13379 1497612 4428115 | Delective CS and VD Odds Ratio | OR 1.01 1.61 0.98 1.47 1.33 1.29 0.96 0.90 0.73 0.96 0.95 | 95%-Cl [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] [1.07; 1.65] [0.91; 1.85] [0.89; 1.03] [0.48; 1.11] [0.85; 1.07] [0.88; 1.02] | Weight 4.3% 4.6% 7.9% 16.3% 33.1% 9.7% 16.9% 15.7% 8.3% 66.9% |
| C Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity. $J^2 = 27\%$, $\tau^2 = 0.0165$, $p = 0.2$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 Black, et al., 2016 Clausen, et al., 2016 Random effects model Heterogeneity. $J^2 = 23\%$, $\tau^2 = 0.0015$, $p = 0.2$ | rude OF Non-Ele T1D 10 18 29 453 5 5 36 725 250 33 336 7 | Cestimate ective CS N 62 38 119 1707 1926 55241 191646 56015 8847 139935 451684 | 236 162 308 7999 1900 9025 1260 68 3762 | N N 1477 452 1248 40529 43706 377268 2286939 252917 13379 1497612 4428115 | Delective CS and VD Odds Ratio | OR 1.01 1.61 0.98 1.47 1.33 1.29 0.96 0.90 0.73 0.96 0.95 | 95%-Cl [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] [1.07; 1.65] [0.89; 1.03] [0.89; 1.03] [0.85; 1.07] [0.88; 1.02] | Weight 4.3% 4.6% 7.9% 16.3% 33.1% 9.7% 16.9% 15.7% 8.3% 66.9% |
| C Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity. $J^2 = 27\%$, $\tau^2 = 0.0165$, $p = 0.2$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 Black, et al., 2016 Random effects model Heterogeneity. $J^2 = 23\%$, $\tau^2 = 0.0015$, $p = 0.2$ Random effects model | rude OF Non-Ele T1D 10 18 29 453 5 5 36 725 250 33 336 7 | Cestimate ective CS N 62 38 119 1707 1926 55241 191646 56015 8847 139935 451684 453610 | 236 162 308 7999 190 9025 1260 68 3762 | N N 1477 452 1248 40529 43706 377268 2286939 252917 13379 1497612 4428115 4471821 | Delective CS and VD Odds Ratio | OR 1.01 1.61 0.98 1.47 1.33 1.29 0.96 0.90 0.73 0.96 0.95 1.05 | 95%-Cl [0.51; 2.02] [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] [1.07; 1.65] [0.89; 1.03] [0.89; 1.03] [0.48; 1.07] [0.88; 1.02] [0.88; 1.02] | Weight 4.3% 4.6% 7.9% 16.3% 33.1% 9.7% 16.9% 15.7% 8.3% 66.9% 10.0% |
| Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity. $J^2 = 27\%$, $\tau^2 = 0.0165$, $p = 0.2$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 Black, et al., 2016 Clausen, et al., 2016 Random effects model Heterogeneity. $J^2 = 23\%$, $\tau^2 = 0.0015$, $p = 0.2$ Random effects model Heterogeneity. $J^2 = 23\%$, $\tau^2 = 0.0015$, $p = 0.2$ | rude OF Non-Ele T1D 10 18 29 453 5 36 725 250 33 336 7 | € estimate ective CS N 62 38 119 1707 1926 55241 191646 56015 8847 139935 451684 453610 | 236 162 308 7999 190 9025 1260 68 3762 | N N 1477 452 1248 40529 43706 377268 2286939 252917 13379 1497612 4428115 4471821 | Delective CS and VD | OR 1.01 1.61 0.98 1.47 1.33 1.29 0.96 0.90 0.73 0.96 0.95 1.05 | 95%-Cl [0.83; 3.13] [0.63; 1.52] [1.32; 1.64] [1.07; 1.65] [0.89; 1.03] [0.78; 1.03] [0.48; 1.11] [0.85; 1.07] [0.89; 1.24] | Weight 4.3% 4.6% 7.9% 16.3% 33.1% 9.7% 16.9% 15.7% 8.3% 16.3% 66.9% |

Figure 2 Summary adjusted OR (aOR) of elective (upper) or non-elective (lower) Caesarean section (CS; N – Total) on offspring Type 1 Diabetes (T1D) compared to vaginal delivery (VD; N – Total).

| Adjusted OR estin | nates comparing Elective C | S and V | /D | |
|---|---|--|--|--|
| Author & year of study / Study design | Odds Ratio | OR | 95%-CI | Weight |
| Study design = case-control | | | | |
| Patterson, et al., 1994 | - | 1.58 | [0.94; 2.67] | 4.0% |
| McKinney, et al., 1997 | | | | 0.0% |
| Robertson & Harrid, 2010 | | | | 0.0% |
| Samuelsson, et al., 2015 | | 0.95 | [0.88; 1.03] | 23.4% |
| Heterogeneity: $I^2 = 72\%$, $\tau^2 = 0.0929$, $p = 0.06$ | | 1.14 | [0.71; 1.85] | 27.4% |
| Study design = cohort | | | | |
| Algert, et al., 2009 | - <u>m</u> | 1.33 | [0.96; 1.85] | 8.2% |
| Khasan, et al., 2014 | | 1.15 | [1.07; 1.24] | 23.8% |
| Black, et al., 2015 | | 1.20 | [0.95; 1.52] | 12.4% |
| Black, et al., 2016 | | 0.71 | [0.49; 1.03] | 6.8% |
| Clausen, et al, 2016 | | 1.10 | [0.99; 1.22] | 21.5% |
| Random effects model Heterogeneity: $I^2 = 47\%$, $\tau^2 = 0.0054$, $p = 0.11$ | | 1.12 | [1.01; 1.24] | 72.6% |
| Random effects model | | 1.09 | [0.97: 1.22] | 100.0% |
| Heterogeneity: $I^2 = 73\%$, $\tau^2 = 0.0127$, $p < 0.01$ | | | | |
| Residual heterogeneity: $I^2 = 55\%$, $p = 0.05$ | 0.5 1 2 | | | |
| | | | | |
| Adjusted OR estimate | es comparing Non-Elective Odds Ratio | CS and OR | 1 VD 95%-Cl | Weight |
| <u>Adjusted OR estimate</u> Author & year of study / Study design | es comparing Non-Elective Odds Ratio | CS and OR | <u>I VD</u> 95%-CI | Weight |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control | es comparing Non-Elective Odds Ratio | CS and OR | <u>I VD</u> 95%-CI | Weight |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1994 | es comparing Non-Elective Odds Ratio | CS and OR 0.91 | 95%-CI | Weight |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 | es comparing Non-Elective Odds Ratio | CS and OR 0.91 | <u>I VD</u> 95%-CI [0.44; 1.87] | Weight 1.6% 0.0% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1997 Robertson & Harrild, 2010 | es comparing Non-Elective Odds Ratio | CS and OR 0.91 | <u>I VD</u> 95%-CI [0.44; 1.87] | Weight 1.6% 0.0% 0.0% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 | Odds Ratio | CS and OR 0.91 1.20 | <u>IVD</u> 95%-CI [0.44; 1.87] [1.06; 1.35] | Weight 1.6% 0.0% 0.0% 22.1% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Huttmeret d ² = 00 × 2 = 0.05 | Odds Ratio | CS and OR 0.91 1.20 1.19 | 95%-CI [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] | 1.6% 0.0% 0.0% 22.1% 23.7% |
| Adjusted OR estimateAuthor & year of study / Study designStudy design = case-controlPatterson, et al., 1994McKinney, et al., 1997Robertson & Harrild, 2010Samuelsson, et al., 2015Random effects modelHeterogeneity: $I^2 = 0\%$, $\tau^2 = 0$, $\rho = 0.46$ | Odds Ratio | CS and OR 0.91 1.20 1.19 | 95%-Cl [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] | Weight 1.6% 0.0% 0.0% 22.1% 23.7% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity. $I^2 = 0\%$, $\tau^2 = 0$, $p = 0.46$ Study design = cohort | Odds Ratio | CS and OR 0.91 1.20 1.19 | 95%-Cl [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] | Ueight 1.6% 0.0% 0.0% 22.1% 23.7% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $I^2 = 0\%$, $\tau^2 = 0$, $p = 0.46$ Study design = cohort Algert, et al., 2009 | Odds Ratio | CS and OR 0.91 1.20 1.19 | 95%-Cl [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] | Weight 1.6% 0.0% 0.0% 22.1% 23.7% 5.9% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $I^2 = 0\%$, $\tau^2 = 0$, $p = 0.46$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 | Odds Ratio | CS and OR 0.91 1.20 1.19 1.31 1.03 | 1 VD 95%-Cl [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] [0.92; 1.87] [0.96; 1.11] | Weight 1.6% 0.0% 22.1% 23.7% 5.9% 28.7% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $I^2 = 0\%$, $\tau^2 = 0$, $p = 0.46$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 Black, et al., 2015 | Odds Ratio | CS and OR 0.91 1.20 1.19 1.31 1.03 0.92 | 1 VD 95%-CI [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] [0.92; 1.87] [0.96; 1.11] [0.76; 1.11] | Weight 1.6% 0.0% 22.1% 23.7% 5.9% 28.7% 14.5% |
| Adjusted OR estimateAuthor & year of study / Study designStudy design = case-controlPatterson, et al., 1997Robertson & Harrild, 2010Samuelsson, et al., 2015Random effects modelHeterogeneity. $I^2 = 0\%, \tau^2 = 0, p = 0.46$ Study design = cohortAlgert, et al., 2014Black, et al., 2015Black, et al., 2016 | Odds Ratio | CS and OR 0.91 1.20 1.19 1.31 1.03 0.92 0.71 | 95%-Cl [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] [0.92; 1.87] [0.96; 1.11] [0.76; 1.11] [0.47; 1.08] | Weight 1.6% 0.0% 22.1% 23.7% 5.9% 28.7% 14.5% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $I^2 = 0\%$, $\tau^2 = 0$, $p = 0.46$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 Black, et al., 2016 Clausen, et al., 2016 | Odds Ratio | CS and OR 0.91 1.20 1.19 1.31 1.03 0.92 0.71 1.00 | 95%-Cl [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] [0.92; 1.87] [0.96; 1.11] [0.76; 1.11] [0.47; 1.08] [0.89; 1.12] | Weight 1.6% 0.0% 0.0% 22.1% 23.7% 5.9% 28.7% 14.5% 4.5% 22.6% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $J^2 = 0\%$, $\tau^2 = 0$, $p = 0.46$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 Black, et al., 2016 Clausen, et al., 2016 Random effects model | Odds Ratio | CS and OR 0.91 1.20 1.19 1.31 1.03 0.92 0.71 1.00 1.00 | 95%-Cl [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] [0.92; 1.87] [0.96; 1.11] [0.76; 1.11] [0.89; 1.12] [0.89; 1.12] [0.92; 1.09] | Weight 1.6% 0.0% 22.1% 23.7% 5.9% 28.7% 14.5% 4.5% 22.6% 26.3% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1994 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $I^2 = 0\%$, $\tau^2 = 0$, $p = 0.46$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 Black, et al., 2016 Clausen, et al., 2016 Clausen, et al., 2016 Random effects model Heterogeneity: $I^2 = 34\%$, $\tau^2 = 0.0032$, $p = 0.19$ | Odds Ratio | CS and OR 0.91 1.20 1.19 1.31 1.03 0.92 0.71 1.00 1.00 | 95%-Cl [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] [0.96; 1.11] [0.76; 1.11] [0.47; 1.08] [0.89; 1.22] [0.92; 1.09] | Weight 1.6% 0.0% 22.1% 23.7% 5.9% 28.7% 14.5% 4.5% 4.5% 76.3% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $I^2 = 0\%$, $\tau^2 = 0$, $p = 0.46$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2014 Black, et al., 2016 Clausen, et al., 2016 Random effects model Heterogeneity: $I^2 = 34\%$, $\tau^2 = 0.0032$, $p = 0.19$ | Odds Ratio | CS and OR 0.91 1.20 1.19 1.31 1.03 0.92 0.71 1.00 1.00 1.00 | 1 VD 95%-Cl [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] [0.92; 1.87] [0.96; 1.11] [0.76; 1.11] [0.89; 1.12] [0.92; 1.09] [0.94; 1.14] | Weight 1.6% 0.0% 22.1% 23.7% 5.9% 28.7% 14.5% 4.5% 22.6% 76.3% 100.0% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1994 McKinney, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $I^2 = 0\%, \tau^2 = 0, p = 0.46$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2015 Black, et al., 2016 Clausen, et al., 2016 Clausen, et al., 2016 Heterogeneity: $I^2 = 34\%, \tau^2 = 0.0032, p = 0.19$ Random effects model Heterogeneity: $I^2 = 52\%, \tau^2 = 0.0087, p = 0.05$ | Odds Ratio | CS and 0.91 1.20 1.19 1.31 1.03 0.92 0.71 1.00 1.00 1.04 | 95%-Cl [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] [0.92; 1.87] [0.96; 1.11] [0.47; 1.08] [0.89; 1.12] [0.92; 1.09] [0.94; 1.14] | Weight 1.6% 0.0% 22.1% 23.7% 5.9% 28.7% 14.5% 4.5% 22.6% 76.3% 100.0% |
| Adjusted OR estimate Author & year of study / Study design Study design = case-control Patterson, et al., 1997 Robertson & Harrild, 2010 Samuelsson, et al., 2015 Random effects model Heterogeneity: $I^2 = 0\%, \tau^2 = 0, p = 0.46$ Study design = cohort Algert, et al., 2009 Khasan, et al., 2015 Black, et al., 2015 Black, et al., 2016 Clausen, et al, 2016 Clausen, et al, 2016 Heterogeneity: $I^2 = 34\%, \tau^2 = 0.0032, p = 0.19$ Random effects model Heterogeneity: $I^2 = 52\%, \tau^2 = 0.0037, p = 0.05$ Residual heterogeneity: $I^2 = 24\%, p = 0.25$ | Odds Ratio | CS and OR 0.91 1.20 1.19 1.31 1.03 0.92 0.71 1.00 1.00 1.04 | 95%-Cl [0.44; 1.87] [1.06; 1.35] [1.06; 1.34] [0.92; 1.87] [0.96; 1.11] [0.76; 1.11] [0.47; 1.08] [0.89; 1.12] [0.92; 1.09] [0.94; 1.14] | Weight 1.6% 0.0% 22.1% 23.7% 5.9% 28.7% 14.5% 4.5% 22.6% 76.3% 100.0% |